

# Yule-Walker estimation of periodic INAR signals

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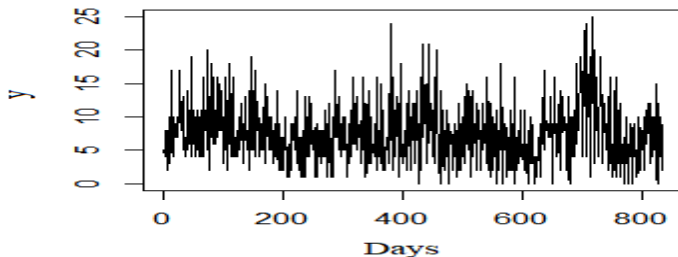
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- 1 Data1: daily number of people who got antibiotics for the treatment of respiratory problems
- 2 Data2: daily number of parcels picked up from one pickup point (PUP)
- 3 Periodic integer-valued autoregressive (PINAR) process
- 4 Yule-Walker estimation
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- 6 Real data applications
- 7 Conclusions

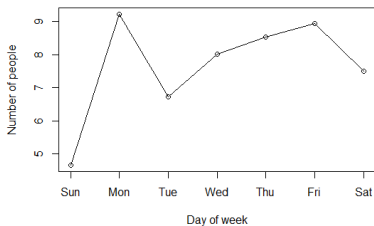
# Data1

- **Time series:** The daily number of people who got antibiotics for treating respiratory problems from the public health care system in the emergency service.
- **Duration:** May 26, 2013, to September 05, 2015, resulting in  $T = 833$  daily ( $n = 119$  weeks) observations.
- **Source:** This real data set was obtained from the network records system welfare of the municipality Vitória-ES, Brazil. (Filho et al., 2021)

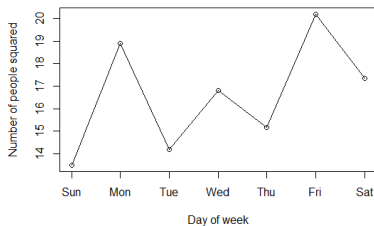


# Basic plots of Data1

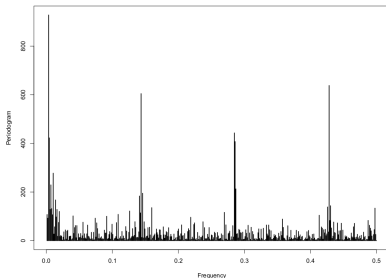
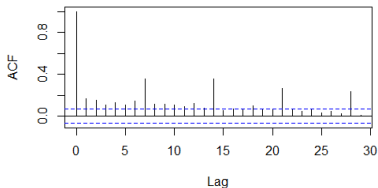
Periodic Mean of Y



Periodic Var of Y



ACF of the Series Y



# Periodic ACF and PACF of Data1

Autocovariance function of periodic time series with period  $S \in \mathbb{N}$ :  $\gamma_s(h) = \text{Cov}(Y_t, Y_{t-h})$ , where  $s = 1, \dots, S$ ,  $h \in \mathbb{N}_0$ , such that  $t \equiv s \pmod{S}$ .

## Periodic ACF

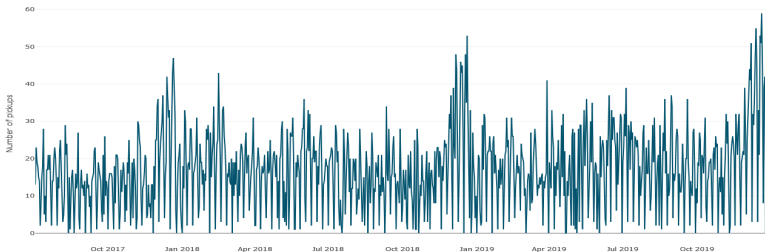
	$h = 1$	$h = 2$	$h = 3$	$h = 4$	$h = 5$	$h = 6$	$h = 7$	$h = 8$	$h = 9$	$h = 10$
Sunday	0.01	<b>0.26</b>	<b>0.18</b>	<b>0.24</b>	<b>0.28</b>	0.11	0.15	0.02	0.07	<b>0.29</b>
Monday	<b>0.38</b>	-0.12	0.14	<b>0.23</b>	<b>0.18</b>	<b>0.19</b>	<b>0.29</b>	0.13	-0.11	0.04
Tuesday	<b>0.33</b>	<b>0.34</b>	-0.02	0.10	<b>0.23</b>	<b>0.39</b>	<b>0.42</b>	<b>0.18</b>	<b>0.37</b>	0.03
Wednesday	<b>0.27</b>	0.05	<b>0.17</b>	0.10	0.16	<b>0.33</b>	<b>0.29</b>	<b>0.23</b>	0.14	0.14
Thursday	<b>0.18</b>	<b>0.36</b>	<b>0.23</b>	<b>0.31</b>	0.01	<b>0.18</b>	<b>0.29</b>	<b>0.22</b>	<b>0.25</b>	0.11
Friday	<b>0.25</b>	0.16	<b>0.20</b>	0.14	0.16	<b>0.17</b>	<b>0.18</b>	<b>0.30</b>	<b>0.23</b>	0.13
Saturday	<b>0.20</b>	0.10	-0.03	-0.05	<b>-0.18</b>	0.03	<b>0.30</b>	0.10	-0.07	0.16

## Periodic PACF

	$h = 1$	$h = 2$	$h = 3$	$h = 4$	$h = 5$	$h = 6$	$h = 7$	$h = 8$	$h = 9$	$h = 10$
Sunday	0.01	<b>0.26</b>	0.12	<b>0.20</b>	<b>0.18</b>	0.00	0.03	-0.02	-0.01	0.16
Monday	<b>0.38</b>	-0.14	0.08	<b>0.18</b>	0.07	0.01	<b>0.22</b>	-0.06	-0.08	-0.02
Tuesday	<b>0.33</b>	<b>0.24</b>	0.00	-0.00	0.15	<b>0.32</b>	<b>0.29</b>	0.01	<b>0.26</b>	0.04
Wednesday	<b>0.27</b>	-0.04	0.10	0.10	0.11	<b>0.27</b>	<b>0.18</b>	0.03	0.09	-0.07
Thursday	<b>0.18</b>	<b>0.33</b>	0.13	<b>0.18</b>	0.01	0.10	<b>0.17</b>	0.04	0.02	-0.02
Friday	<b>0.25</b>	0.12	0.10	0.06	0.03	<b>0.18</b>	0.08	<b>0.18</b>	0.13	-0.05
Saturday	<b>0.20</b>	0.05	-0.07	-0.11	<b>-0.21</b>	0.08	<b>0.26</b>	0.03	-0.13	<b>0.21</b>

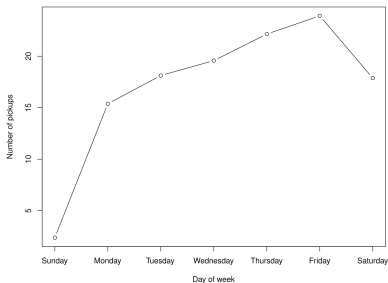
bold: significant correlation

- **Time series:** The daily number of parcels picked up from one pickup point (PUP) at a PUP management company.
- **Duration:** July 3, 2017 to December 29, 2019, resulting in  $T = 910$  daily ( $n = 130$  weeks) observations.
- **Source:** <https://github.com/cabani/ForecastingParcels> (Nguyen, et al., 2023)

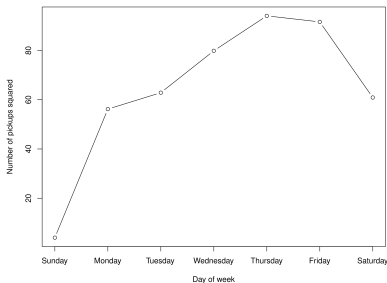


# Basic plots of Data2

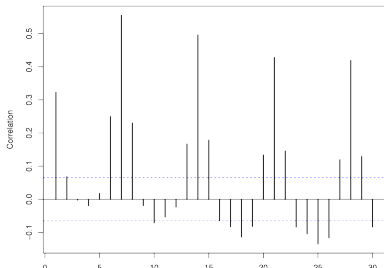
Periodic Mean



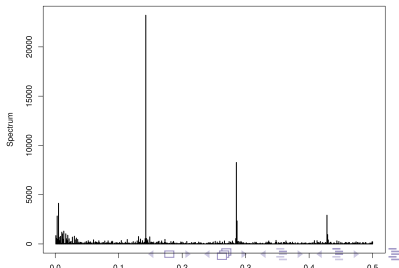
Periodic Variance



ACF of Pickup series



Periodogram of Pickup series



# Periodic ACF and PACF of Data2

## Periodic ACF

	$h = 1$	$h = 2$	$h = 3$	$h = 4$	$h = 5$	$h = 6$	$h = 7$	$h = 8$	$h = 9$	$h = 10$
Sunday	0.072	0.008	0.118	-0.021	0.036	0.120	-0.042	0.000	-0.058	-0.012
Monday	<b>0.261</b>	<b>0.215</b>	<b>0.370</b>	<b>0.287</b>	<b>0.321</b>	0.075	0.169	0.084	0.184	0.186
Tuesday	<b>0.328</b>	<b>0.438</b>	<b>0.241</b>	<b>0.208</b>	0.081	<b>0.281</b>	0.060	0.168	<b>0.205</b>	0.115
Wednesday	<b>0.548</b>	<b>0.479</b>	<b>0.373</b>	<b>0.215</b>	<b>0.342</b>	0.171	<b>0.222</b>	<b>0.238</b>	<b>0.238</b>	<b>0.232</b>
Thursday	<b>0.486</b>	<b>0.450</b>	<b>0.196</b>	<b>0.278</b>	<b>0.196</b>	<b>0.222</b>	<b>0.308</b>	<b>0.406</b>	<b>0.245</b>	0.096
Friday	<b>0.521</b>	0.149	<b>0.381</b>	<b>0.351</b>	<b>0.314</b>	<b>0.398</b>	<b>0.368</b>	<b>0.363</b>	0.097	<b>0.337</b>
Saturday	<b>0.244</b>	<b>0.332</b>	<b>0.238</b>	<b>0.341</b>	<b>0.312</b>	<b>0.443</b>	<b>0.406</b>	<b>0.260</b>	<b>0.234</b>	0.135

## Periodic PACF

	$h = 1$	$h = 2$	$h = 3$	$h = 4$	$h = 5$	$h = 6$	$h = 7$	$h = 8$	$h = 9$	$h = 10$
Sunday	0.072	-0.011	<b>0.174</b>	<b>0.237</b>	0.009	-0.037	-0.068	-0.045	-0.088	-0.021
Monday	<b>0.261</b>	0.142	0.114	<b>0.240</b>	0.091	0.015	0.033	0.081	0.067	-0.022
Tuesday	<b>0.328</b>	<b>0.326</b>	<b>0.260</b>	-0.116	0.085	0.142	0.118	-0.045	0.027	0.005
Wednesday	<b>0.548</b>	<b>0.290</b>	0.002	0.134	-0.004	<b>0.297</b>	0.065	<b>0.204</b>	-0.086	-0.156
Thursday	<b>0.486</b>	<b>0.264</b>	0.064	-0.008	0.152	0.117	0.152	0.063	-0.104	0.017
Friday	<b>0.521</b>	0.027	0.096	0.109	-0.021	-0.025	-0.081	<b>0.202</b>	-0.080	0.010
Saturday	<b>0.244</b>	<b>0.327</b>	<b>0.258</b>	0.074	0.164	0.157	0.003	-0.052	0.020	0.158

bold: significant correlation

## Observations:

- periodicity (varying environment)
- strong seasonality



# Overviews on count time series

- Davis RA, Fokianos K, Holan SH, Joe H, Livsey J, Lund R, Pipiras V, Ravishanker N. (2021) Count time series: a methodological review. *Journal of the American Statistical Association* 116:1533–1547.
- Davis RA, Holan SH, Lund R, Ravishanker N. (2016) *Handbook of Discrete-Valued Time Series* CRC Press, New York, NY.
- Fokianos, K. (2012) Count time series models. *Handbook of Statistics.*, C. R. Rao, C. Rao, and V. Govindaraju (eds); 30, 315–347, Elsevier, Amsterdam.
- Weiss CH. (2018) *An Introduction to Discrete-Valued Time Series*. John Wiley & Sons, Boca Raton, FL.

# Branching process with immigration (BPI)

$$X_k = \sum_{j=1}^{X_{k-1}} \xi_{k,j} + \varepsilon_k, \quad k \in \mathbb{N}, \quad X_0 = 0$$

$\xi_k := \{\xi_{k,j} \mid j \in \mathbb{N}\}$  and  $\{\varepsilon_k \mid k \in \mathbb{N}\}$  i.i.d. sequences of r.v.'s  
To avoid degeneracy:  $P(\varepsilon = 0) < 1$

Reformulation by (generalized) thinning operator:

$$X_k = \xi_k \circ X_{k-1} + \varepsilon_k, \quad k \in \mathbb{N}, \quad X_0 = 0$$

**Parameters:**  $m := E[\xi]$ ,  $\sigma^2 := \text{Var}[\xi]$ ,  $\lambda := E[\varepsilon]$ ,  $b^2 := \text{Var}[\varepsilon]$

**Existence:** always since  $k \in \mathbb{N}_0$

**Classification:**

$m < 1$	$m = 1$	$m > 1$
<b>subcritical</b>	<b>critical</b>	<b>supercritical</b>

# Integer valued autoregression (INAR)

**INAR(1) model**, Al-Osh and Alzaid (1987), McKenzie (1985)

$$Y_k = \alpha_k \circ Y_{k-1} + \varepsilon_k, \quad k \in \mathbb{Z}$$

$\{\alpha_k \circ \mid k \in \mathbb{Z}\}$ , are i.i.d. **binomial thinning operators** with parameter  $\alpha \in [0, 1]$ ,  $\{\varepsilon_k \mid k \in \mathbb{Z}\}$  are i.i.d.  $\mathbb{N}_0$ -valued r.v.'s  
To avoid degeneracy:  $P(\varepsilon = 0) < 1$

Reformulation in BPI form:

$$Y_k = \sum_{j=1}^{Y_{k-1}} \xi_{k,j} + \varepsilon_k, \quad k \in \mathbb{Z}$$

$\{\xi_{k,j} \mid k \in \mathbb{Z}, j \in \mathbb{N}\}$  i.i.d. **Bernoulli** r.v.'s with mean  $\alpha$

**Parameters:**  $\alpha := E[\xi]$ ,  $\lambda := E[\varepsilon]$ ,  $b^2 := \text{Var}[\varepsilon]$

**Existence:**  $\alpha < 1$  (subcritical)

In the critical case  $\alpha = 1$ , there is no solution!

# Periodic INAR signals

$\{Y_t\}$  is called **periodic INAR (PINAR) process** with period  $S \in \{2, 3, \dots\}$  and autoregressive orders  $\mathbf{p} := \{p_s\}$  if

$$Y_{kS+s} = \sum_{i=1}^{p_s} \alpha_{s,i}^k \circ Y_{kS+s-i} + \varepsilon_{kS+s},$$

$k \in \mathbb{Z}$ , where  $\{\alpha_{s,i}^k\}$  are i.i.d. binomial thinning operators with mean  $\alpha_{s,i} \geq 0$  for all  $i = 1, \dots, p_s, s = 1, \dots, S$

$Y_{kS+s}$ : the value of series during the  $s$ th season of period  $k$

**Autoregressive** parameters:  $\{\alpha_{s,i}, | i = 1, \dots, p_s, s = 1, \dots, S\}$

**Immigration process**  $\{\varepsilon_t\}$ : periodic sequence of  $\mathbb{N}_0$ -valued r.v.'s, i.e., for each  $s$ ,  $\{\varepsilon_{kS+s} | k \in \mathbb{Z}\}$  are i.i.d.r.v.'s

**Immigration** parameters:  $\lambda_s = \mathbf{E}(\varepsilon_{kS+s}) \in \mathbb{R}_+$ , and

$\sigma_s^2 = \text{Var}(\varepsilon_{kS+s}) > 0, k \in \mathbb{Z}$ .

Periodic AR (PAR) model:

$$X_{kS+s} = \sum_{i=1}^{p_s} \alpha_{s,i} X_{kS+s-i} + \varepsilon_{kS+s}$$

# Periodic time series models

Examples:

- **PINAR(1) model**, Monteiro et al. (2010)
- **INAR(1)<sub>S</sub> model**, Bourguignon, Vasconcellos, Reisen, I (2014), Buteikis and Leipus (2020)
- **PINAR(2) model** with periodic immigration, Morina et al. (2011)
- **PINAR(1, S) model**, Filho et al. and I (2021, 2024)
- **PINARMA(p,q) model** Bentarzi and Aries (2020)

Periodic models (**PARMA**): Gladyshev (1961), Gardner et al. (2006), and Hurd and Miamee (2007). Inference: Lund and Basawa (2000) and Basawa and Lund (2001).

Seasonal models (**SARMA**): Chatfield and Prothero (1973)

**SPARIMA** models: Basawa, Lund, and Shao (2004), Koopman et al. (2006), Hindrayanto et al. (2010)

# BP(I) in varying environment

General inhomogeneous branching processes:

- Domain of peer-to-peer file sharing networks, Adar and Huberman (2000), Zhao et al. (2005)
- Modeling biodiversity or macroevolution, Aldous and Popovic (2005), Haccou and Iwasa (1996)
- Epidemic-type Aftershock Sequence (ETAS) in seismology, Farrington et al. (2003)

Heterogeneous INAR models (Bernoulli offspring):

- Understanding and predicting consumers' buying behavior, Böckenholt (1999)
- Modeling the premium in the bonus-malus scheme of car insurance, Gourieroux and Jasiak (2004)

The supercritical case was studied by Goettge (1976), Cohn and Hering (1983), Jagers and Nerman (1985), D'Souza and Biggins (1992, 1993), D'Souza (1994).

# State-space (MBPI) representation

**Main tool:** stationary vector representation á la Gladyshev  
Define the **matricial thinning operator**  $\Xi \circ = (\xi_{i,j} \circ)$ , Latour (1997), as

$$(\Xi \circ \mathbf{Y})_i = \sum_{j=1}^S \xi_{i,j} \circ Y_j, \quad i = 1, \dots, S$$

Let  $\mathbf{Y}_k = (Y_{kS+1}, \dots, Y_{kS+S})^\top$  and  $\varepsilon_k = (\varepsilon_{kS+1}, \dots, \varepsilon_{kS+S})^\top$   
Implicit state-space representation by **VINAR<sub>S</sub>(p) model**:

$$\mathbf{Y}_k = \sum_{i=0}^p A_i^k \circ \mathbf{Y}_{k-i} + \varepsilon_k$$

where  $\{A_i^k \circ\}$  are  $S \times S$  matricial binomial thinning operators defined as  $A_i^k \circ = (a_{r,s}^{k,i} \circ)$ ,  $i = 0, 1, \dots, p$ , where  $a_{r,s}^{k,0} \circ := \alpha_{r,r-s}^k \circ$  if  $r > s$  and  $0 \circ$  otherwise (i.e., if  $r \leq s$ ), and  $a_{r,s}^{k,i} \circ := \alpha_{r,iS+r-s}^k \circ$  for all  $i = 1, \dots, p$  and  $r, s = 1, \dots, S$ , where  $\alpha_{s,i}^k \circ := 0 \circ$  if  $i > p_s$  for all  $k \in \mathbb{Z}$ .

**Assumptions.** Suppose that  $E\|\varepsilon_0\| < \infty$  and  $\rho(A_0 + A_1 + \dots + A_p) < 1$ , where  $A_i$  denotes the mean matrix of  $A_i^k \circ$ ,  $i = 0, 1, \dots, p$ .

The **VINAR<sub>S</sub>(p)** model has a unique non-anticipative, strictly stationary solution  $\{\mathbf{Y}_k\}$  with finite mean vector. The  $\mathbb{N}_0^S$ -valued stochastic process  $\{\mathbf{Y}_k\}$  is  $\{\mathcal{G}_k\}$ -adapted, ergodic  $p$ th order homogeneous Markov chain.

The **PINAR<sub>S</sub>(p)** model has a unique non-anticipative, periodically strictly stationary solution  $\{Y_t\}$  with finite mean and period  $S$ . The  $\mathbb{N}_0$ -valued stochastic process  $\{Y_t\}$  is  $\{\mathcal{H}_t\}$ -adapted, periodic inhomogeneous Markov chain with orders  $p_s$ ,  $s = 1, \dots, S$ .



# Examples

Consider the  $\text{PINAR}(1)_S$  model. The **characteristic polynomial** of this model simplifies to  $P(z) = z^S - \prod_{j=1}^S \alpha_j$  and the condition  $\rho(A_0 + A_1) < 1$  is equivalent to  $\prod_{j=1}^S \alpha_j < 1$ .

Consider the case  $S = 2$ , with  $p_1 = p_2 = 2$ , i.e., the  $\text{PINAR}_2(2, 2)$  model. Then

$$A_0 = \begin{bmatrix} 0 & 0 \\ \alpha_{2,1} & 0 \end{bmatrix}, \quad A_1 = \begin{bmatrix} \alpha_{1,2} & \alpha_{1,1} \\ 0 & \alpha_{2,2} \end{bmatrix}, \quad A_0 + A_1 = \begin{bmatrix} \alpha_{1,2} & \alpha_{1,1} \\ \alpha_{2,1} & \alpha_{2,2} \end{bmatrix}$$

The **characteristic polynomial** is given by

$$P(z) = (z - \alpha_{1,2})(z - \alpha_{2,2}) - \alpha_{1,1}\alpha_{2,1}$$

It can be seen that  $\alpha_{1,2} + \alpha_{2,2} - \alpha_{1,2}\alpha_{2,2} + \alpha_{1,1}\alpha_{2,1} < 1$  is a necessary and sufficient stationarity condition. Note that this condition can be rewritten as  $\alpha_{1,1}\alpha_{2,1} < (1 - \alpha_{1,2})(1 - \alpha_{2,2})$ . See Darolles et al. (2019)

# Yule-Walker estimation

Parameters of  $VINAR_S(p)$ :  $A_i$ 's, and  $\lambda = E(\varepsilon)$ ,  $\Lambda = \text{Var}(\varepsilon)$

First-order or Markovian state-space representation results in

$VINAR_{Sp}(1)$  model with parameters:  $A_0, A_1, \lambda$ , and  $\Lambda$

Let  $\mu = E(\mathbf{Y})$ ,  $\Sigma = \text{Var}(\mathbf{Y}) = \Gamma(0)$  and  $\Gamma = \Gamma(1)$

We have the Yule-Walker equations:

$$\mu = (I - A_0 - A_1)^{-1} \lambda$$

$$\Sigma - \Gamma \Sigma^{-1} \Gamma^T = (I - A_0)^{-1} D \left( (I - A_0)^{-1} \right)^T$$

$$\Gamma \Sigma^{-1} = (I - A_0)^{-1} A_1$$

where  $D := \Lambda + \text{diag}((V_0 + V_1)\mu)$  and  $V_0, V_1$  are the Bernoulli variance of  $A_0, A_1$ . Number of equations:  $(3/2)S(S + 1)$

**Scenario S1:** the immigration is uncorrelated within a period  
2nd equation: Cholesky factorization of a Schur complement

**Scenario S2:** the immigration is correlated within a period,  
the autoregressive orders are fixed  $S$

3rd equation: LU factorization if  $A_1$  is upper triangular

# Result of the YW estimation for Pickup data under S1

col	1	2	3	4	5	6	7
$\hat{A}_0$	0	0	0	0	0	0	0
	0.164	0	0	0	0	0	0
	0.029	0.270	0	0	0	0	0
	0.224	0.253	0.437	0	0	0	0
	0.096	-0.058	0.281	0.213	0	0	0
	0.072	-0.026	0.012	0.163	0.275	0	0
	-0.022	-0.003	0.024	0.016	-0.014	0.041	0
$\hat{A}_1$	-0.006	0.15	0.138	-0.037	0.165	0.155	-0.141
	-0.042	-0.065	-0.032	0.05	0.232	0.068	-0.226
	0.063	0.078	0.034	-0.034	0.078	0.221	0.204
	0.002	0.025	-0.05	0.113	0.113	-0.001	-0.548
	0.127	-0.173	-0.072	0.247	0.014	0.234	-0.361
	0.122	0.121	0.013	-0.151	0.065	0.126	0.306
	-0.002	0.018	0.035	-0.017	-0.028	0.066	-0.131
$\hat{\lambda}$	4.501	10.72	5.623	2.159	6.648	1.261	0.712

# Result of the YW estimation for Pickup data under S2

col	1	2	3	4	5	6	7
$\hat{A}_1$ + $\hat{A}_0$	0.165	0.114	0.009	-0.067	0.151	0.261	0.085
	0.882	-0.22	-0.057	0.316	-0.014	0.119	-0.523
	0.58	-0.703	0.098	0.183	0.161	0.25	-0.505
	0.509	-0.503	-0.254	0.26	0.307	0.249	-0.889
	0.315	0.057	-0.024	0.067	0.285	0.148	-0.161
	0.371	0.005	0.252	0.003	0.208	0.093	0.192
$\hat{\lambda}$	-0.24	0.027	0.083	-0.159	-0.213	0.655	-0.312
$\hat{\Lambda}$	3.581	2.095	10.309	12.971	17.621	-0.009	1.6176
	34.79	-31.62	-17.22	-4.83	20.22	-11.55	8.56
	-31.62	75.94	64.14	53.66	-13.3	6.73	-6.25
	-17.22	64.14	129.02	94.83	16.87	-6.22	-4.84
	-4.83	53.66	94.83	137.49	22.87	2.47	2.16
	20.22	-13.3	16.87	22.87	69.41	-3.85	5.6
	-11.55	6.73	-6.22	2.47	-3.85	29.58	-23.68
	8.56	-6.25	-4.84	2.16	5.6	-23.68	29.1

# Conclusions

- Real count time series are presented which possess periodicity and seasonality.
- A periodic INAR model is proposed.
- YW method under two scenarios estimates the model parameters.
- Monte Carlo simulations verify the effectiveness of the proposed methodology.
- The proposed model is successfully fitted to real data.

# Thank you for your attention! <sup>1</sup>

## References:

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